



## Numerical Likelihood Analysis of Cosmic Ray Anisotropies

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### Abstract

A numerical likelihood approach to the determination of cosmic ray anisotropies is presented which offers many advantages over other approaches. It allows a wide range of statistically meaningful hypotheses to be compared even when full sky coverage is unavailable, can be readily extended in order to include measurement errors, and makes maximum unbiased use of all available information.

### 1. Introduction

The search for anisotropies in the cosmic ray distribution (and of other objects – see for example [4,5]) is of increasing interest with the advent of the Pierre Auger Observatory (for a review of the physics up to this time, see [1]). The issues involved are, however, subtle and complicated by limited statistics at the highest energies and nonuniform sky coverage. The first full-sky search for UHECR anisotropies using a standard approach is presented in this conference [2]. This paper presents a maximum likelihood approach which is well-adapted to further studies in anticipation of much larger data sets.

### 2. Expansion in Spherical Harmonics

It is convenient to define a set of real spherical harmonics  $\{\psi_{\ell,m}\}$  for  $\ell = 0, 1, 2, \dots$  and  $m = -\ell, -(\ell-1), \dots, 0, \dots, (\ell-1), \ell$  by

$$\psi_{\ell,m}(\theta, \phi) = \begin{cases} k_{\ell}^{|m|} P_{\ell}^{|m|}(\cos \theta) \cos(m\phi) & \text{for } m = -\ell, \dots, -1, \\ k_{\ell}^0 P_{\ell}^0(\cos \theta) & \text{for } m = 0, \\ k_{\ell}^m P_{\ell}^m(\cos \theta) \sin(m\phi) & \text{for } m = 1, \dots, \ell \end{cases} \quad (1)$$

which are orthonormal with respect to the usual measure  $\sin(\theta)d\theta d\phi$  and integration over  $\theta$  from 0 to  $\pi$  and  $\phi$  from 0 to  $2\pi$ , the  $P_{\ell}^m$  are associated Legendre polynomials of the first kind, and the normalization constants  $k_m^{\ell}$  are:

$$k_0^0 = \frac{1}{\sqrt{2\pi}}, k_{\ell}^0 = \sqrt{\frac{2\ell+1}{4\pi}}, \text{ and } k_{\ell}^m = \sqrt{\frac{2\ell+1}{2\pi} \frac{(\ell-m)!}{(\ell+m)!}} \quad (2)$$

The natural measure of anisotropy for a spherical distribution is in terms of these spherical harmonics, as each  $\ell$  labels, in a coordinate-independent fashion, just how much of each irreducible  $SO(3)$  representation is present.

In a perfect world with infinite statistics and complete sky coverage there are now many possible approaches to estimating how much of each of these components is present in a distribution, or, better, what is the likelihood that a given function  $f(\theta, \phi)$  with Fourier-Legendre expansion  $f(\theta, \phi) = \sum_{\ell=0}^{\infty} \sum_{m=-\ell}^{\ell} a_{\ell,m} \psi_{\ell,m}$  representing the probability density of sources gives rise to the observed distribution  $g(\theta, \phi)$ . The coefficients can be extracted from the usual integral

$$a_{\ell,m} = \int_0^{2\pi} \int_0^{\pi} f(\theta, \phi) \psi_{\ell,m}(\theta, \phi) \sin \theta d\theta d\phi \quad (3)$$

Of course this gives no measure of what sort of error should be associated with the determined values of each coefficient. Alternative approaches are to fit for the coefficients by minimizing some  $\chi^2$ -like quantity like

$$\int_0^{2\pi} \int_0^{\pi} \frac{\left| g(\theta, \phi) - \sum_{\ell=0}^{\infty} \sum_{m=-\ell}^{\ell} a_{\ell,m} \psi_{\ell,m}(\theta, \phi) \right|^2}{\sigma^2} \sin \theta d\theta d\phi \quad (4)$$

with  $\sigma$  a suitable measure of error, or to compute and maximize a corresponding likelihood that the hypothesized distribution parametrized by the  $a_{\ell,m}$  gives rise to the the observed distribution  $g$ . Of course this is statistically unreasonable for a finite number of sources (*i.e.* to provide an infinite number of coefficients!), In general a decision must be made to truncate the expansion at some value of  $\ell$  to make the sum finite, but now the general phenomenon of aliasing risks that, for example, a fit allowing for  $\ell = 0, 1$  might give misleading results for observed data which is drawn from a purely  $\ell = 2$  distribution, say.

A more serious problem is that should part of the sky be unobserved, there is now no way to calculate anything! This is not a trivial point. An attempt to find anisotropies based only on observations in the Northern hemisphere with zeroes inserted for the whole Southern hemisphere would be wildly in error if some simple extrapolation were made to the unobserved region of the sky – especially if there were something bright and as-yet undetected in the South! The real challenge is to say something statistically meaningful with the data that is actually available. Clearly, observing only the Northern hemisphere and seeing a good degree of isotropy should increase one’s net belief in overall isotropy of the full sky, while leaving open the possibility of a staggeringly bright or empty sky in the South. One approach is to try to make functions which would be orthonormal over the observed region of the sky, but it’s not clear that this has much physical meaning as it elevates lack of acceptance to a status comparable to the  $SO(3)$  invariance of space embodied in the  $\psi_{\ell,m}$ . The following is a proposal for what seems to make good statistical sense.

### 3. A Likelihood Proposal

Based on the above observations, we make the following proposal: keep the spherical harmonics as always with the (necessarily) truncated Fourier-Legendre expansion and construct an unbinned likelihood [3] function in which one clearly specifies which values of  $\ell$  are included in the sum. An unbinned likelihood function, as described below, automatically makes maximum use of all detected information, allows for measurement errors to be included easily, and is easy to implement numerically. Most importantly, however, the likelihood is *not* to be normalized as it stands. Rather, we take the Bayesian approach to likelihood which says that likelihoods give us ways to update our prior experimental data or guesses in light of new information. This will mean that we are able to present results on various hypotheses about data without bias, and with the easy inclusion of other data from the same, or other experiments.

To be concrete, we specialize here to the case where  $g$  is a sum of delta functions representing sources  $i$  of unit intensity at  $\theta_i, \phi_i$  and later discuss how one can treat the case of these being at uncertain locations, or taking into account other properties such as intensity, energy, composition, *etc.*. These will appear as natural generalizations to the approach described.

The (unnormalized!) likelihood  $L(\{\ell\}|\theta_i, \phi_i)$  that the measured  $\theta_i, \phi_i$  arise from  $f(\theta, \phi) = \sum_{\ell}^{\infty} \sum_{m=-\ell}^{\ell} a_{\ell,m} \psi_{\ell,m}$  where the sum over  $\ell$  is specified according to whatever hypothesis is being tested (*i.e.* just taking  $\{\ell\} = \{0, 1\}$  allows for uniform and dipole contributions and no others, while  $\{\ell\} = \{2\}$  would be pure quadrupole) is

$$L(\{\ell\}|\theta_i, \phi_i) = \frac{f(\theta_i, \phi_i)}{\int \prod_{\ell} \prod_{m=-\ell}^{\ell} da_{\ell,m} f(\theta, \phi)} \quad (5)$$

where the integral in the denominator is over all the parameters that can vary and over the range in which the parameters are allowed to vary. An important caveat is that for  $f$  to represent a sensible probability distribution it should never be negative, and this must be checked. Two approaches are possible: one is to restrict the domain over which coefficients range so that the function is strictly positive (this is not actually very difficult in practice since distributions are often nearly uniform with small fluctuations superimposed) or to take as a probability distribution function some positive function of  $f$  in place of  $f$  above. In astronomy [4] it is not uncommon to see  $\exp(f)$ . The choice ultimately represents the unavoidable presence of some (often hidden) assumption about what a sensible prior (*i.e.* in the absence of data) is for the likelihood – a point to which we return later.

By construction then  $L(\{\ell\}|\theta_i, \phi_i)$  is normalized so that its integral over all the parameters that can vary (the coefficients  $a_{\ell,m}$  included in the truncated Fourier-Legendre expansion, and the total likelihood  $L_{TOT}$ , which is a function of

those  $a_{\ell,m}$  is

$$L_{TOT}(\{\ell\}) = \prod_i L(\{\ell\}|\theta_i, \phi_i) \quad (6)$$

This is a relative likelihood, and while no absolute normalization is possible (nor should it be!) if part of the sky is unobserved, it is now very useful for two types of calculation. If one has a prior expectation for the distribution of the  $a_{\ell,m}$  (which might be that they are all a priori equally likely) then  $L_{TOT}$  can be multiplied by this and the product treated as a normalized likelihood distribution for the  $a_{\ell,m}$  themselves. This is, of course, potentially dangerous, but does allow one to see how the new data (the measured  $(\theta_i, \phi_i)$ ) should cause one to revise earlier beliefs. Such a likelihood can be maximized with respect to the  $a_{\ell,m}$  and if not enough data is present to test the hypothesis (for example, more coefficients to fit for than data points) the fit will respond by just not converging (that is to say, there will be flat directions in the likelihood as a function of the parameters meaning that one can't decide) - the beauty of this approach is that it is, by construction, correct. When results are obtained, one automatically gets the values of the parameters, and their entire likelihood distributions from which errors (which need not even be Gaussian) can be extracted. One can even do things like fix some parameters to those given by a favoured theory and then repeat the fitting and then obtain likelihoods for the correctness of that theory.

More objectively, one can compute relative likelihoods in which the prior drops out, so that it is reasonable to ask (even in the absence of full sky coverage!) what the relative likelihood  $L_{REL}$  is of pure dipole distribution to one admitting uniform, dipole and quadrupole components:  $L_{REL} = \frac{L_{TOT}(\{1\})}{L_{TOT}(\{0,1,2\})}$ . If one wants statements made to be about all energies over  $x$ , then one just uses the data points with energies over  $x$ . Similarly, data can be selected by composition and (even relative) anisotropies be searched for as functions of composition, energy, time, *etc.* Extensions to uncertainty in direction are trivial to include: simply divide a given event into a large number,  $M$ , say of subevents distributed appropriately and count each in the likelihood with a weight  $1/M$  - this numerically convolves this uncertainty with the likelihood, and the size of any additional errors introduced by the procedure can be studied by varying  $M$ .

#### 4. References

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